

12 Feb. 2009

(1) Lecture # 11

To see whether a single sample  $(y_1, \dots, y_n)$  from a pop. w/ a mean  $\mu$  supports a theory that  $\mu = \mu_0$  (theoretical value), build a 95% CI for  $\mu$  using the t machinery  $\hat{=}$  see if  $\mu$  is in the interval; if not, data don't support theory @ 95% CI; if so, data do support theory at that level.

" When  $\mu_0$  is not in 95% interval, people say difference between theory  $\mu_0$   $\hat{=}$  data ( $\bar{y}$ ) is statistically significant (statsig); when  $\mu_0$  is not in 99% CI, the difference is highly statsig.

Q: is difference between  $\mu_0$   $\hat{=}$   $\bar{y}$  practically significant (practsig)?

ex. Like asking whether a difference of  $0.7^\circ\text{C}$  ( $25^\circ - 24.3^\circ$ ) is large in biological terms (from  $\mu_0$  crab case study from lecture #10)

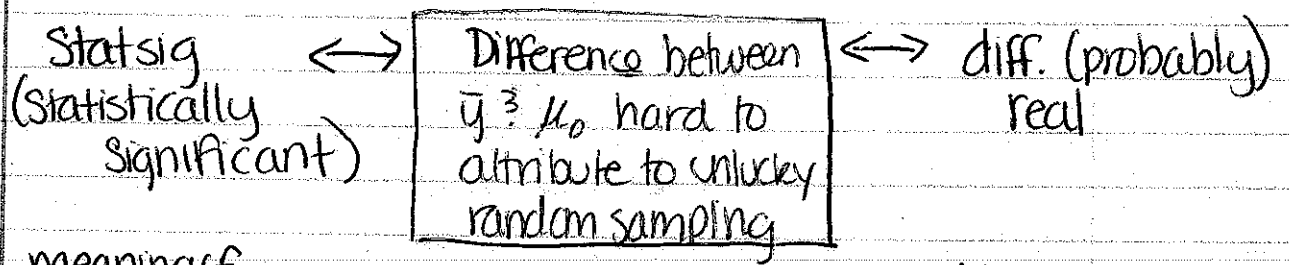
" Don't know - but of important

" Practical significance is the 1st question

Q: - Is a difference that is statsig real?

Not real: easy to attribute to unlucky random sampling (if in 95% CI?)

Real: hard to attribute to unlucky random sampling = statsig



meaning of CIs

Neyman 1925

(2) Lecture #11

- Back to Crab Case Study:

95% CI for  $\mu$  was (24.5°C, 25.6°C)

Does this mean that

$P_f (24.5^\circ C < \mu < 25.6^\circ C) = 95\%$

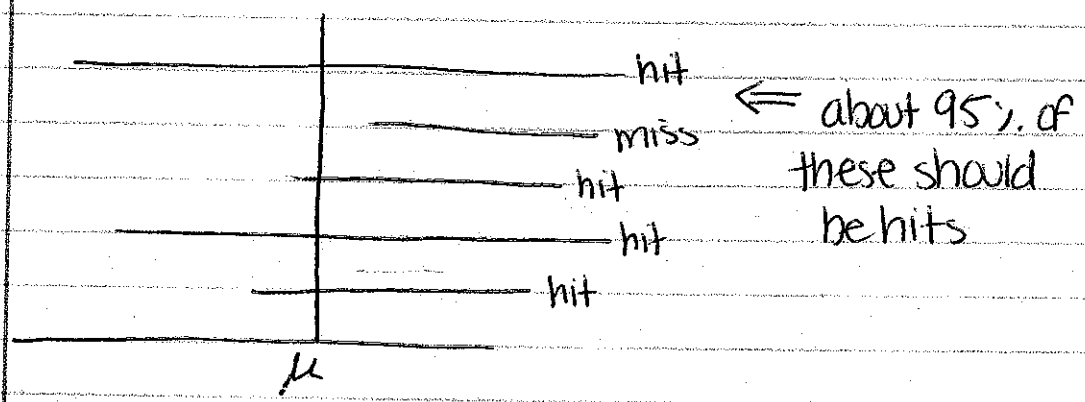
relative frequency

In other words,

confidence = probability

No!  $\mu$  is a Fixed unknown constant that's not changing across repetitions

Q. OK, what does the interval mean?



Can't tell if any single CI you build is a hit or miss, because we don't know  $\mu$ ; so our confidence in CI's is in the process of building the intervals, not in any single interval

Case Study Lab Rats

Food

lab rats: 12 animals, 10 turn left to food

IL [1's] n=12  
OR [0's]

mean  $\frac{10}{12} = 83\%$

Is the difference between  $\frac{10}{12} = 83\% \hat{p}$  50% large?

- large in practical terms?  $\Rightarrow$  huge

- large in statistical terms?  $\Rightarrow$

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(3) Lecture # 11

ask idologists

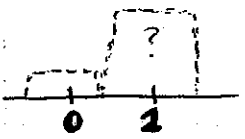
population  
All similar lab rats (relevant)

left?  
N=? (big)  
 $\begin{bmatrix} 1^s \\ 3^s \\ 0^s \end{bmatrix}$

mean  $\mu = p = ?$

SD  $\sigma = ?$

pop. hist.



sample  
the observed lab rats

$y=1$   
 $n=0$

left?

$\begin{bmatrix} 1^s \\ 3^s \\ 0^s \end{bmatrix} n=12$

mean  $\bar{y} = \hat{p} = \frac{10}{12}$

"p hat"  $= .83 = 83\%$

(actual) like SRS IID

$\begin{bmatrix} \phantom{1^s} \\ \phantom{3^s} \\ \phantom{0^s} \end{bmatrix} n=12$

mean  $\hat{p} = ?$

ex:  $\frac{9}{12} = 75\%$

Imag. Data

possible  $\hat{p}^s$

$\begin{bmatrix} \hat{p}^s \\ 83\% \\ 75\% \\ \vdots \end{bmatrix} M = \infty$

long run mean:

expected value of  $\hat{p} = p$

EV of  $\hat{p} = E_{IID}(\hat{p}) =$

$E_{IID}(\bar{y}) = \mu = p$

$E(\hat{p}) = p$  math fact

CI:  $y \pm t_{n-1}^{.95} \hat{SE}_{IID}(\bar{y})$

"continuous outcome"

approximate:  $\hat{p} \pm 1.96 \hat{SE}_{IID}(\hat{p})$

0-1 outcomes  $83\% \pm 2(11\%)$

long run SD:

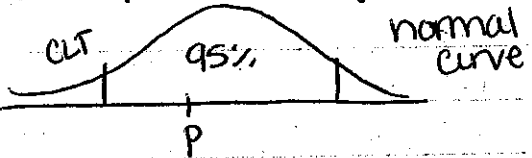
est. standard error of  $\hat{p} = 11\%$

$\hat{SE}$  of  $\hat{p} = \hat{SE}_{IID}(\hat{p}) =$  (estimated)

$SE_{IID}(\bar{y}) = \sigma/\sqrt{n}$

long run histogram of  $\hat{p} =$

$\bar{y}$  if n is large



earlier math fact: if the pop has only 2 values in it

$\sigma = \left( \begin{matrix} \text{larger value} \\ \text{smaller value} \end{matrix} \right) \sqrt{\left( \begin{matrix} \text{proportion of} \\ \text{larger values} \end{matrix} \right) \left( \begin{matrix} \text{proportion of} \\ \text{smaller values} \end{matrix} \right)}$

So here; larger value = 1  
smaller value = 0

$\therefore$  SD of a 0-1 pop is

$\sigma = \sqrt{p(1-p)}$  math fact with 0-1 pop.

So:  $SE_{IID}(\hat{p}) = \frac{\sqrt{p(1-p)}}{\sqrt{n}} = \sqrt{\frac{p(1-p)}{n}}$

\* is not in 95% CI, so the difference between 83% and 50% is statistically significant (is probably real)

$= \sqrt{\frac{(0.83)(0.07)}{12}} = 0.108 = 10.8\% = 11\%$

(4) Lecture #11

Inferential Summary

Pop.	Unknown quantity of interest	$p = \text{pop \% of lab animals that would turn left (toward food)}$
Sample	estimate of $p$	$\hat{p} = 83\%$
imag. data set	give or take for $\hat{p}$ as est of $p$	$\hat{SE}(\hat{p}) = 11\%$
imag. data set	95% CI for $p$	$\hat{p} \pm 1.96 \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = (61\%, 100\%)$ truncated at 0% $\approx$ 100%

CI<sup>s</sup> are only 1 way to do statistical inference - the other is called hypothesis tests (Neyman & Pearson, 1930<sup>s</sup>)  
& significance tests (Fisher 1930<sup>s</sup>)

Draper will argue hyp./sig. tests are not as good as CI<sup>s</sup> but we need to know both.

- ① null hypothesis ( $H_0$ ) |  $\mu = 24.3^\circ\text{C}$  | "theory correct"  
 ② alternative hypothesis ( $H_A$ ) |  $\mu \neq 24.3^\circ\text{C}$  | "theory wrong"

- ① The difference between  $\mu_0 = 24.3^\circ\text{C}$  &  $\bar{y} = 25.0^\circ\text{C}$  is due to unlucky random sampling (this is a logical possibility)  
 ② No, the difference between 24.3 & 25.0 is real

Neyman's logic: try (null) on for size, see if discrepancy (Fisher's) between

\* (how data came out) vs (how data should have come out if null true) \*\*  
 is large, if yes, favor alt ("reject null"), if not, favor null ("fail to reject null")

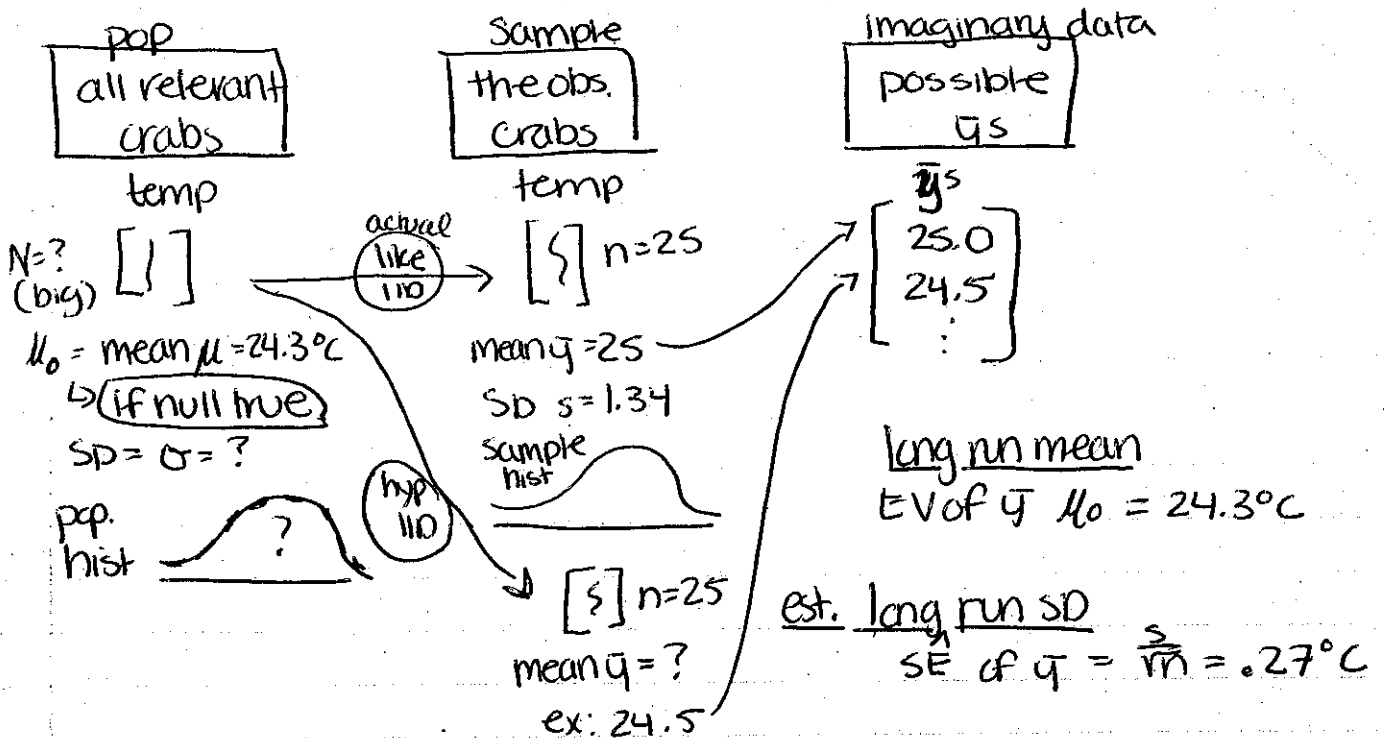
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\* (how data came out):  $\bar{y} = 25.0^\circ\text{C}$  [VS]

\*\* (how data should have come out if null true)

$E_{\text{IID}}(\bar{y}) \text{ (if null true)} = \mu_0 = 24.3^\circ\text{C}$



how measure discrepancy?

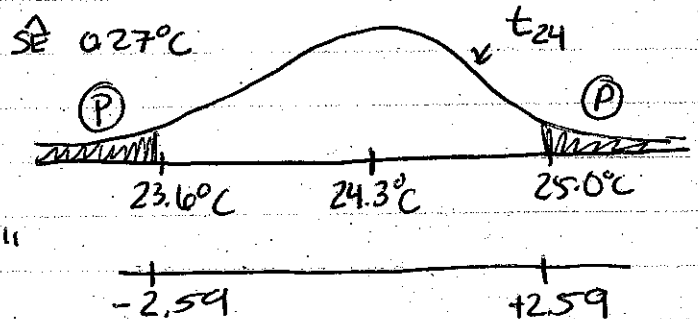
$\frac{s}{\sqrt{n}} \rightarrow \left( \text{SE of } \bar{y} \text{ if null is true} \right) = \frac{25^\circ\text{C} - 24.3^\circ\text{C}}{0.27^\circ\text{C}}$

t test  $\frac{+0.7^\circ\text{C}}{0.27^\circ\text{C}} = \frac{\text{"signal"}}{\text{"noise"}} =$

$+2.59 = t$  "the t statistic"

long run hist if null true accounting for uncertainty in  $\sigma$

$\frac{25.0^\circ - 24.3^\circ}{0.27^\circ} = +2.59 = t$



Chance if null true, of getting data as extreme as, or more extreme than, what I got = numerical surprise = p value measure